# PREVENTION OF CARDIOVASCULAR DISEASE AT POPULATION LEVEL 

Modelling strategies for primary prevention of cardiovascular disease

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## West Midlands Health Technology Assessment Collaboration

The West Midlands Health Technology Assessment Collaboration (WMHTAC) is an organisation involving several universities and academic groups who collaboratively undertake research synthesis to produce health technology assessments. Most of our members are based in the Department of Public Health, Epidemiology \& Biostatistics, University of Birmingham, however other members are drawn from a wide field of expertise including economists and mathematical modellers from the Health Economics Facility, University of Birmingham.

WMHTAC produce systematic reviews, health technology assessments and economic evaluations for NHS R\&D HTA programme (NCCHTA), the National Institute for Health and Clinical Excellence (NICE), and for the health service in the West Midlands. WMHTAC also undertakes methodological research on research synthesis, and provides training in systematic reviews and health technology assessment.

## Name of other institution(s) involved

WMHTAC work in close collaboration with the Peninsula Technology Appraisal Group (PenTAG) with respect to providing support to the CPHE.

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## Executive Summary

The National Institute for Health and Clinical Excellence ('NICE' or 'the Institute') has been asked by the Department of Health (DH) to develop guidance on a public health programme aimed at preventing cardiovascular disease (CVD) in different populations.

This report is the economic modelling report to be delivered to the Programme Development Group (PDG). It complements three effectiveness reports and an economic review addressing question 1 defined in the final scope as:

Which multiple risk-factor interventions are effective and cost effective in the primary prevention of CVD within a given population? Where the data allows, how does the effectiveness and cost effectiveness of interventions vary between different population groups?

At the request of the programme development group (PDG), the scope of the modelling was extended beyond multiple risk-factor programmes for which there is direct evidence of effectiveness to consider also single risk factor programmes for which there is direct or indirect evidence of effectiveness. An example of this is modelling of a legislative programme to reduce the use of trans fatty acids (TFAs) in food. Here the effectiveness can be estimated based on the known relative risks incurred by the consumption of TFAs.

A spreadsheet model has been developed which will allow a relative risk to be applied to each year's risk of primary CVD within the population. An alternative form allows percentage reductions in cholesterol and systolic blood pressure to be applied separately for males and females. In their current forms, the models have been built on the assumption that these effects apply uniformly across age and risk groups and, in the case of the "Relative Risk" model, across the ten years and equally for males and females. It would not be difficult to amend the model to allow variation in the effect by any of these factors if such amendment were felt appropriate.

The model has been applied to estimate the effects in terms of outcomes such as quality adjusted life years (QALYs) gained and savings in health care costs for a
given effectiveness. This gives an estimate of how much it would be worth spending to achieve such an outcome.

The results strongly suggest that any legislative intervention that is likely to achieve an appreciable reduction in risk of CVD can be expected to produce a net cost saving to the public sector as well as improving health. Only if a very large sum of money needs to be spent in implementing the legislation would this cease to be the case.

No attempt has been made to include the effects of smoking cessation in the analysis of multi-factor interventions using the "Risk Factor Modifying" model. To do so with any attempt at realism would require the proportion of smokers in each of the risk groups used for the modelling.

Similarly, the analysis is restricted to effects on primary CVD prevention. An intervention which is recommended on the basis of this analysis and is known to have only beneficial effects on other aspects of health can be recommended more strongly as a result.

## 1. Introduction

The National Institute for Health and Clinical Excellence ('NICE' or 'the Institute') has been asked by the Department of Health (DH) to develop guidance on a public health programme aimed at preventing cardiovascular disease (CVD) in different populations.

This report is the economic modelling report to be delivered to the Programme Development Group (PDG). It complements three effectiveness reports and an economic review addressing question 1 defined in the final scope as:

Which multiple risk-factor interventions are effective and cost effective in the primary prevention of CVD within a given population? Where the data allows, how does the effectiveness and cost effectiveness of interventions vary between different population groups?

The PDG subsequently responded to stakeholder feedback and explicitly also considered single risk factor interventions. Accordingly, this report also considers such interventions.

### 1.1 Background

Evaluating complex changes between populations is problematic for a number of reasons, for example: it is difficult to design studies which evaluate entire cities, regions or countries; control sites can become 'contaminated' (that is, if the intervention affects people living in the control area); unreasonable expectations about the speed of effect; and failure to address 'upstream' influences such as policy or manufacturing practices. Some population programmes have been accompanied by a substantial reduction in the rate of CVD deaths. However, the degree to which these are attributable to the programme is debatable.

The precise nature of the populations and interventions to be covered, and those which are not included are defined in the final scope (as extended by the PDG) as follows:

| POPULATION |  |
| :--- | :--- |
| COVERED BY GUIDANCE | NOT COVERED BY GUIDANCE |
| Groups to be covered are populations <br> defined on a geographical basis. The area <br> will usually be at least a region of a <br> country (such as Merseyside) or an urban <br> or rural area (such as Paisley and <br> Nottingham or New Forest). In the UK, <br> the geographical area would not be less <br> individuals who are clinically diagnosed <br> as being at high risk of developing - or <br> than what is currently covered by a | CVD. However, as populations include <br> Primary Care Trust. A population could <br> also be made up of people living in a <br> designated geographical area that fulfils <br> the criteria above who also share a specific <br> characteristic, such as all South Asian men |
| will have some relevance for them. |  |
| (Individuals at high risk of developing |  |
| over 50 who live in Sheffield. Populations |  |
| will include both adults and children. |  |$\quad$| CVD are covered by other NICE |
| :--- |
| guidance, see section 6.) |


|  |  |
| :--- | :--- |
| OR Programmes that include a <br> pharmacological element alongside a <br> broader, non-pharmacological multiple <br> risk-factor approach (as indicated in <br> 4.2.1a) will be included when they involve <br> a primary prevention element and where <br> data can be disaggregated to allow <br> consideration of the impact of the non- <br> pharmacological elements. | OR Interventions which focus on <br> screening for CVD risk factors (for <br> example, cholesterol-level screening) <br> and do not attempt to modify them |
| OR Natural experiments, such as changes <br> in the diet of Eastern Europeans brought <br> about by social change, where relevant <br> evidence is available |  |

### 1.2 Structure of this report

The structure of this report is as follows:

- Chapter 2 describes the interventions to be modelled and the sources of information for the modelling.
- Chapter 3 describes the modelling process.
- Chapter 4 discusses the review findings, highlighting their applicability, limitations and any gaps.


## 2. Interventions and information sources

The cost-effectiveness review (Andronis et al, 2009) identified a number of potentially cost-effective community based interventions. However, none of the studies found in that review could be regarded as directly applicable, for two main reasons. First, the rules of evaluation (such as costing perspective, and discounting rates to be applied) did not correspond with those currently required by NICE: this could be resolved by repeating the analysis using the appropriate evaluation rules. Second, the background against which the evaluation was carried out may have changed considerably since the time of the intervention. For example, an important part of the HeartBeat Wales programme was the introduction of food labelling. Since such food labelling is now widespread, the benefit of that part of the programme would already be included in the background for any new programme. To produce a useful estimate of the effect of such a programme now would require detailed information on the changes in the background to the programme, together with the difference such changes would make to the effectiveness of the programme.

At the request of the programme development group (PDG), and in recognition of consistent feedback from consultation with stakeholders, the scope of the modelling was extended beyond multiple risk factor programmes to consider also single risk factor programmes for which there is direct or indirect evidence of effectiveness. An example of this is modelling of a legislative programme to reduce the use of trans fatty acids (TFAs) in food. Here the effectiveness can be estimated based on the known relative risks incurred by the consumption of TFAs.

Finally, it is possible to model the effects in terms of outcomes such as quality adjusted life years (QALYs) gained and savings in health care costs for a given effectiveness. This gives an estimate of how much it would be worth spending to achieve such an outcome.

### 2.1 Information sources

The information required for the modelling consists of three parts. First is the background pattern of risks. For this we have used the Joint British Societies' Report
(Joint British Societies, 2005) for the risk factor equation and the distribution of risk factors in the population. We have used national statistics for other information about the general population.

The second type of information relates to the effectiveness of potential interventions. For this, we have used the results of our reviews together with expert papers presented to the PDG.

Finally, we need to be able to convert cases prevented into outcomes such as QALYs gained and healthcare costs saved. For this, we have relied heavily on the inputs to previous modelling undertaken for NICE (Ward et al, 2005).

## 3. Details of modelling undertaken

Specific interventions have been modelled in accordance with their potential effects. Additionally, a range of hypothetical analyses have been carried out to assess the possible effects of hypothetical interventions. The aim in each case is to estimate the following outcomes:

- CVD cases prevented or postponed;
- CVD deaths prevented or postponed;
- Life-years gained;
- Quality adjusted life-years (QALYs) gained;
- Cost savings to the NHS resulting from cases prevented or postponed.

In the case of specific interventions where it is possible to assess the costs of providing the intervention, a full cost-effectiveness analysis has been completed. Where this has not been possible, an estimate has been made of the maximum cost for such an intervention to be cost-effective at thresholds of $£ 20,000$ and $£ 30,000$ per QALY gained.

### 3.1 The basic model structure

The modelling process consists of five stages:

1. Determining the estimated outcomes as far as QALYs lost and costs to the NHS for a case of CVD;
2. Assessing the pattern of CVD cases prevented or postponed for an intervention of known effectiveness, applied to a single combination of age, sex, and risk;
3. Combining the results from stages 1 and 2 to estimate the potential outcomes for a single combination of age, sex, and risk;
4. Aggregating the results from stage 3 across all levels of risk to estimate the potential outcomes for a single combination of age and sex;
5. Aggregating the results from stage 4 to give total estimated outcomes at population level.

At stage 1, a lifetime horizon is feasible. However, at stage 2, it is only appropriate to apply a time horizon of around 10 years, given the nature of the risk equation and the
assumptions necessary. Accordingly the model gives an estimate of lifetime effects from a reduction in the number of cases within 10 years. Lifetime benefits would clearly be greater.

### 3.1.1 Determining the estimated outcomes for a case of CVD

The first stage of the model consists of an estimate of the expected lifetime costs, life years and QALYs following a first CVD event. Comparing these to life expectancy without an event gives us the life year loss and QALY loss from such an event. The main source of information to answer this question is the report by ScHARR (Ward et al, 2005) which considered the use of statins for the prevention of CVD.

Table 3.1 CVD event types and additional mortality for a 65-year-old male

| Event type | Proportion | Additional mortality |  |
| :--- | :---: | :---: | :---: |
|  |  | First year | Later yrs |
| Stable Angina | 0.214 | 0.0070 | 0.0070 |
| Unstable Angina | 0.083 | 0.1077 | 0.0124 |
| Myocardial Infarction | 0.173 | 0.0626 | 0.0159 |
| Fatal CHD | 0.097 |  |  |
| TIA | 0.100 | 0.0348 | 0.0348 |
| Stroke | 0.270 | 0.0520 | 0.0208 |
| Fatal CVD (ex CHD) | 0.063 |  |  |

Sources: Ward et al (2005), pages 138, 142.

Consider the example of a 65 year old male having a first CVD event. Available data from the ScHARR statins model (Ward et al, 2005) are summarised in Tables 3.1 and 3.2. Costs following an event have been re-estimated where possible, and otherwise inflated to 2008 prices. There is some methodological disagreement on how to account for co-morbidities in handling quality of life scores. The simplest approach would be to credit CVD-free individuals with full health, and apply the quality of life scores in Table 3.2 following a CVD event. However, such an approach would be contrary to the general principle of conservative modelling that has been taken in this work.. An alternative, used in the base case analysis by ScHARR, is to apply population norms for quality of life scores, and treat the values in Table 3.2 as
multipliers applied to these population norms. The ScHARR report (Ward et al, 2005, p. 151) quotes the results of a regression analysis giving a baseline utility of $1.06-0.004 n$ at age $n$ years. For simplicity, the utility at the age of event has been applied to calculate QALYs lost as a result of an event. This somewhat offsets the fact that the population utilities include some CVD patients, but is still likely to be a conservative valuation overall.

Table 3.2 CVD event types - costs and quality of life effects for a 65 -year-old male

| Event type | Proportion | Costs following event |  | QoL following event |
| :--- | :---: | :---: | :---: | :---: |
|  |  | First year | Later yrs |  |
| Stable Angina | 0.214 | $£ 232$ | $£ 232$ | 0.808 |
| Unstable Angina | 0.083 | $£ 541$ | $£ 232$ | 0.770 |
| Myocardial Infarction | 0.173 | $£ 5,244$ | $£ 232$ | 0.760 |
| Fatal CHD | 0.097 | $£ 1,341$ |  |  |
| TIA | 0.100 | $£ 1,224$ | $£ 304$ | 1.000 |
| Stroke | 0.270 | $£ 9,259$ | $£ 2,489$ | 0.629 |
| Fatal CVD (ex CHD) | 0.063 | $£ 8,102$ |  |  |

Sources: Ward et al (2005), pages 146, 153. Costs following event have been inflated to 2008, except where GP contact costs have been identified, in which case the updated PSSRU cost has been used.

Data from the Government Actuary's Department (2009) indicate a life expectancy for 65 year old males of 17.29 years. The age-dependent utility is 0.784 , giving a qualityadjusted life expectancy of $0.784 \times 17.29=13.56$ QALY. For such an individual with stable angina, the estimated life expectancy is reduced to 15.42 years (see Appendix 1 for method used here). The quality-adjusted life expectancy is then calculated as $0.784 \times 0.808 \times 15.42=9.77$ QALY. Thus the undiscounted life year loss and QALY loss for the event are $17.29-15.42=1.87$ and $13.56-9.77=3.79$ respectively. Further, the lifetime cost estimate for such an event is $15.42 \times £ 232=£ 3,578$.

Applying a discount rate of $3.5 \%$, the estimated discounted life expectancy is 10.84 years before the event, reduced to 10.08 years following the event. The discounted quality-adjusted life expectancy before and after the event are thus respectively
$0.784 \times 10.84=8.50$ and $0.784 \times 0.808 \times 10.08=6.38$ QALY. This gives a discounted loss of 0.76 life years ( 2.12 QALY), with lifetime costs of $10.08 \times £ 232=£ 2,338$.

Similar calculations are made for the other types of CVD event: the results for a 65 year old male are shown in Tables 3.3 and 3.4. The overall results for all ages from 40 to 90 male and female are given in Appendix 2.

Table 3.3 CVD events - undiscounted effects for a 65-year-old male

| Event type | Proportion | LY lost | QALY lost | Lifetime costs |
| :--- | :---: | :---: | :---: | :---: |
| Stable Angina | 0.214 | 1.87 | 3.79 | 3,578 |
| Unstable Angina | 0.083 | 4.41 | 5.78 | 3,297 |
| Myocardial Infarction | 0.173 | 4.36 | 5.85 | 8,011 |
| Fatal CHD | 0.097 | 17.29 | 13.56 | 1,341 |
| TIA | 0.100 | 6.50 | 5.10 | 4,202 |
| Stroke | 0.270 | 4.97 | 7.48 | 37,434 |
| Fatal CVD (ex CHD) | 0.063 | 17.29 | 13.56 | 8,102 |
| Overall | 1 | 6.28 | 7.00 | 13,593 |

Table 3.4 CVD events - discounted effects for a 65-year-old male

| Event type | Proportion | LY lost | QALY lost | Lifetime costs |
| :--- | :---: | :---: | :---: | :---: |
| Stable Angina | 0.214 | 0.76 | 2.12 | 2,338 |
| Unstable Angina | 0.083 | 1.91 | 3.11 | 2,380 |
| Myocardial Infarction | 0.173 | 1.89 | 3.17 | 7,088 |
| Fatal CHD | 0.097 | 10.84 | 8.50 | 1,341 |
| TIA | 0.100 | 2.97 | 2.33 | 3,313 |
| Stroke | 0.270 | 2.19 | 4.16 | 28,307 |
| Fatal CVD (ex CHD) | 0.063 | 10.84 | 8.50 | 8,102 |
| Overall | 1 | 3.27 | 4.00 | 10,539 |

### 3.1.2 Assessing the pattern of cases prevented or postponed for an intervention of

 known effectivenessTwo versions of the model have been created. One uses a relative risk applied to the annual risk of a first CVD event, while the other uses modifications to the risk factor
equation directly. For convenience, the relative risk version of the model is described in detail, and the variation for the other version follows.

Table 3.5 shows the calculations for a 65 -year-old male with a 10 -year CVD risk of $12.5 \%$. The intervention is assumed to have a relative risk of 0.9 for each of the 10 years modelled.

Table 3.5 CVD events over 10 years for 65 -year-old male with $12.5 \%$ 10-year risk

|  |  | no intervention |  |  | intervention |  | undiscounted |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| discounted |  |  |  |  |  |  |  |
| Age | OC death | risk CVD | CVD free | risk CVD | CVD free | cas prev | cas prev |
| A | B | C | D | E | F | G | H |
| 65 | 0.00791 | 0.00819 | 0.98396 | 0.00737 | 0.98477 | 0.00081 | 0.00081 |
| 66 | 0.00844 | 0.01074 | 0.96518 | 0.00967 | 0.96703 | 0.00103 | 0.00100 |
| 67 | 0.00911 | 0.01197 | 0.94493 | 0.01077 | 0.94789 | 0.00111 | 0.00103 |
| 68 | 0.00989 | 0.01285 | 0.92356 | 0.01156 | 0.92766 | 0.00114 | 0.00103 |
| 69 | 0.01052 | 0.01354 | 0.90147 | 0.01219 | 0.90671 | 0.00115 | 0.00100 |
| 70 | 0.01121 | 0.01412 | 0.87878 | 0.01271 | 0.88515 | 0.00113 | 0.00095 |
| 71 | 0.01242 | 0.01462 | 0.85518 | 0.01316 | 0.86266 | 0.00111 | 0.00090 |
| 72 | 0.01341 | 0.01506 | 0.83100 | 0.01356 | 0.83955 | 0.00107 | 0.00084 |
| 73 | 0.01428 | 0.01546 | 0.80647 | 0.01392 | 0.81605 | 0.00103 | 0.00078 |
| 74 | 0.01553 | 0.01582 | 0.78138 | 0.01424 | 0.79194 | 0.00097 | 0.00071 |
|  |  |  |  |  | totals | 0.01055 | 0.00906 |

Column A simply shows the age at the start of each year. Column B gives the risk of other causes death in the year, defined as the probability of other causes death conditional on survival to the start of the year. Yearly all-cause death rates have been adjusted for the proportion of CVD deaths to obtain this estimate. Column C is the assumed risk profile in the absence of an intervention, so that (for example) someone who has survived to age 70 without a CVD event has a probability of 0.01412 of a first CVD event in the next year. This is obtained from applying the algorithm of Anderson et al (1991). Column D gives the probability of CVD-free survival to the end of each year. It is calculated by a formula such as $D_{67}=\left(-B_{67}\right)-C_{67} D_{66}$, where $D_{67}$ represents the entry in column $D$ on the row for age 67 , and so on. In other words, the CVD-free survival is multiplied each year by two factors, one representing other cause death and a second representing first CVD event. The multiplicative formula allows appropriately for competing risks. Column E is simply calculated as $E_{65}=0.9 C_{65}$, and so on, showing the intervention effect. Then column F is calculated as $F_{67}=$ 《 $-B_{67} \boldsymbol{\nearrow}-E_{67} F_{66}$, by analogy with column D. Next, the estimated cases
prevented each year are calculated by a formula such as
$G_{67}=\boldsymbol{Q}_{66}-D_{67} \doteq \boldsymbol{\epsilon}_{66}-F_{67} \doteq$ Finally, column H gives the discounted figure from column $G$ at a rate of $3.5 \%$ as required by NICE. For example, $H_{67}=G_{67} /\left(035^{2}\right.$;

For the alternative risk-factor modification version of the model, the annual risk of CVD in column $E$ is calculated directly from the modified risk equation instead of working in terms of column C .

### 3.1.3 Estimating outcomes for a single combination of age, sex, and risk

The age and sex specific outcomes as illustrated in Section 3.1.1 can be combined with the estimated pattern of events saved as calculated in Section 3.1.2. Multiplying the cases prevented from column $G$ in Table 3.5 by the undiscounted results in Appendix 2 gives us the results in Table 3.6 for a 65 -year-old male with a $12.5 \% 10$ year CVD risk. For discounted results, the relevant figures in Appendix 2 are discounted to the age at event. Multiplying by the figures in column H thus gives us outcomes discounted to the starting age as required. These are shown in Table 3.7.

Table 3.6 Undiscounted outcomes for 65-year-old male with a 12.5\% 10-year CVD risk

| Age | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | Cost saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 65 | 0.00081 | 0.00013 | 0.0051 | 0.0057 | 11 |
| 66 | 0.00103 | 0.00017 | 0.0061 | 0.0068 | 14 |
| 67 | 0.00111 | 0.00018 | 0.0062 | 0.0069 | 14 |
| 68 | 0.00114 | 0.00018 | 0.0059 | 0.0067 | 14 |
| 69 | 0.00115 | 0.00018 | 0.0056 | 0.0063 | 14 |
| 70 | 0.00113 | 0.00016 | 0.0063 | 0.0066 | 14 |
| 71 | 0.00111 | 0.00016 | 0.0058 | 0.0060 | 13 |
| 72 | 0.00107 | 0.00015 | 0.0052 | 0.0054 | 12 |
| 73 | 0.00103 | 0.00015 | 0.0046 | 0.0049 | 12 |
| 74 | 0.00097 | 0.00014 | 0.0041 | 0.0043 | 11 |
| Totals | 0.01055 | 0.00160 | 0.0549 | 0.0596 | 129 |

Table 3.7 Discounted outcomes for 65 -year-old male with a $12.5 \% 10$-year CVD risk

| Age | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | Cost saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 65 | 0.00081 | 0.00013 | 0.0027 | 0.0032 | 9 |
| 66 | 0.00100 | 0.00016 | 0.0031 | 0.0038 | 10 |
| 67 | 0.00103 | 0.00017 | 0.0031 | 0.0038 | 11 |
| 68 | 0.00103 | 0.00016 | 0.0030 | 0.0036 | 10 |
| 69 | 0.00100 | 0.00016 | 0.0028 | 0.0034 | 10 |
| 70 | 0.00095 | 0.00014 | 0.0031 | 0.0035 | 10 |
| 71 | 0.00090 | 0.00013 | 0.0028 | 0.0031 | 9 |
| 72 | 0.00084 | 0.00012 | 0.0025 | 0.0028 | 8 |
| 73 | 0.00078 | 0.00011 | 0.0022 | 0.0025 | 8 |
| 74 | 0.00071 | 0.00010 | 0.0019 | 0.0021 | 7 |
| Totals | 0.00906 | 0.00138 | 0.0271 | 0.0319 | 92 |

### 3.1.4 Aggregating across different risk factors

To aggregate across different risk factors requires the distribution of risk factors in any given age group. The best source of this information is the JBS 2 report (Joint British Societies, 2005, page v11).This gives prevalence in 10 year age groups. For ease of modelling, the "under $10 \%$ " risk group has been taken at $7.5 \%$ and the "over $30 \%$ " risk group at $32.5 \%$. The undiscounted and discounted results are shown for an intervention with relative risk 0.9 in Tables 3.8 and 3.9 respectively.

Table 3.8 Undiscounted results for 65-year-old males

| 10-year risk | proportion | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | Cost saved |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.075 | 0.010 | 0.00648 | 0.00098 | 0.0334 | 0.0361 | 79 |
| 0.125 | 0.088 | 0.01055 | 0.00160 | 0.0549 | 0.0596 | 129 |
| 0.175 | 0.197 | 0.01443 | 0.00220 | 0.0759 | 0.0825 | 178 |
| 0.225 | 0.223 | 0.01806 | 0.00276 | 0.0958 | 0.1045 | 224 |
| 0.275 | 0.153 | 0.02145 | 0.00329 | 0.1149 | 0.1255 | 267 |
| 0.325 | 0.329 | 0.02460 | 0.00379 | 0.1330 | 0.1457 | 308 |
| Overall |  | 0.01924 | 0.00295 | 0.1028 | 0.1123 | 239 |

Table 3.9 Discounted results for 65 -year-old males

| 10-year risk | proportion | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | Cost saved |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.075 | 0.010 | 0.00551 | 0.00083 | 0.0164 | 0.0193 | 56 |
| 0.125 | 0.088 | 0.00906 | 0.00138 | 0.0271 | 0.0319 | 92 |
| 0.175 | 0.197 | 0.01251 | 0.00191 | 0.0376 | 0.0444 | 127 |
| 0.225 | 0.223 | 0.01579 | 0.00242 | 0.0477 | 0.0564 | 160 |
| 0.275 | 0.153 | 0.01892 | 0.00291 | 0.0573 | 0.0681 | 192 |
| 0.325 | 0.329 | 0.02189 | 0.00339 | 0.0666 | 0.0793 | 223 |
| Overall |  | 0.01694 | 0.00261 | 0.0513 | 0.0608 | 172 |

Tables 3.10 and 3.11 respectively show the overall figures undiscounted and discounted for age groups between 40 and 79 both males and females.

Table 3.10 Undiscounted estimate of effects for an intervention with relative risk 0.9

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 9.01 | 1.07 | 67.9 | 96.5 | 157 |
| Males 50-59 | 14.21 | 2.08 | 99.6 | 120.1 | 216 |
| Males 60-69 | 19.24 | 2.95 | 102.8 | 112.3 | 239 |
| Males 70-79 | 20.46 | 2.88 | 65.3 | 68.0 | 195 |
| Females 40-49 | 7.56 | 0.75 | 61.2 | 88.0 | 158 |
| Females 50-59 | 9.16 | 1.29 | 75.8 | 91.0 | 162 |
| Females 60-69 | 12.50 | 2.02 | 84.9 | 90.0 | 180 |
| Females 70-79 | 15.24 | 2.29 | 64.7 | 63.8 | 164 |

Table 3.11 Discounted estimate of effects for an intervention with relative risk 0.9

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 7.67 | 0.90 | 21.5 | 36.0 | 82 |
| Males 50-59 | 12.30 | 1.79 | 39.5 | 53.4 | 133 |
| Males 60-69 | 16.94 | 2.61 | 51.3 | 60.8 | 172 |
| Males 70-79 | 18.24 | 2.58 | 40.5 | 44.6 | 158 |
| Females 40-49 | 6.40 | 0.63 | 17.3 | 30.3 | 77 |
| Females 50-59 | 7.82 | 1.08 | 27.1 | 37.4 | 95 |
| Females 60-69 | 10.81 | 1.76 | 39.1 | 45.5 | 124 |
| Females 70-79 | 13.40 | 2.01 | 37.7 | 39.5 | 130 |

### 3.1.5 Aggregating to obtain population effects

To obtain population effects, the size of the relevant population must be taken into account. The relevant sources here are ONS data (Office of National Statistics, 2009) for the total population by 10 -year age groups in England and Wales. The ScHARR report on statins (Ward et al, 2005, p. 140) was used to provide an estimate of the prevalence of CVD history within the population, thereby allowing an estimate of the total population for primary prevention. The results are shown in Table 3.12. Scaling up the results from Tables 3.10 and 3.11 gives us total estimates as shown in Tables 3.13 and 3.14 respectively.

Table 3.12 Population estimates

|  |  | CVD prevalence per | CVD free population in |
| :---: | :---: | :---: | :---: |
| Males 40-49 | Population in 000s | 1000 | 000 s |
| Males 50-59 | 3927 | 7.2 | 3898 |
| Males 60-69 | 3257 | 23.2 | 3181 |
| Males 70-79 | 2659 | 36.1 | 2563 |
| Females 40-49 | 1748 | 44.2 | 1671 |
| Females 50-59 | 3996 | 3.04 | 3983 |
| Females 60-69 | 3346 | 11.0 | 3309 |
| Females 70-79 | 2815 | 21.4 | 2754 |

Table 3.13 Undiscounted estimates of total population effects for an intervention with relative risk 0.9

|  | 000s of cases prevented | 000s of deaths prevented | $\begin{aligned} & \text { 000s of LY } \\ & \text { gained } \end{aligned}$ | 000s of QALYs gained | £millions saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 35 | 4.2 | 265 | 376 | 612 |
| Males 50-59 | 45 | 6.6 | 317 | 382 | 688 |
| Males 60-69 | 49 | 7.6 | 264 | 288 | 614 |
| Males 70-79 | 34 | 4.8 | 109 | 114 | 325 |
| Females 40-49 | 30 | 3.0 | 244 | 351 | 630 |
| Females 50-59 | 30 | 4.3 | 251 | 301 | 536 |
| Females 60-69 | 34 | 5.6 | 234 | 248 | 495 |
| Females 70-79 | 31 | 4.6 | 130 | 129 | 331 |
| Totals | 289 | 40.6 | 1814 | 2188 | 4232 |

Table 3.14 Discounted estimates of total population effects for an intervention with relative risk 0.9

|  | 000s of cases <br> prevented | 000 s of deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 30 | 3.5 | 84 | 140 | 319 |
| Males 50-59 | 39 | 5.7 | 126 | 170 | 423 |
| Males 60-69 | 43 | 6.7 | 131 | 156 | 441 |
| Males 70-79 | 30 | 4.3 | 68 | 75 | 264 |
| Females 40-49 | 26 | 2.5 | 69 | 121 | 307 |
| Females 50-59 | 26 | 3.6 | 90 | 124 | 314 |
| Females 60-69 | 30 | 4.8 | 108 | 125 | 343 |
| Females 70-79 | 27 | 4.1 | 76 | 80 | 262 |
| Totals | 251 | 35.2 | 751 | 990 | 2671 |

### 3.2 Results for specific interventions

Specific interventions modelled here are multifactor interventions based on the North Karelia and Stanford Five City projects and the possible effects of legislation to ban transfats and reduce salt consumption.

### 3.2.1 The North Karelia project

From our previous review, the effect of the North Karelia project included net percentage reductions in serum cholesterol of $3 \%$ for men and $1 \%$ for women, and systolic blood pressure of $3 \%$ for men and $5 \%$ for women. Including only those effects in the "Risk Equation Modifying" model gives undiscounted results by age groups as shown in Table 3.15 and discounted results in Table 3.16. The combined population results are given in Tables 3.17 and 3.18.

Table 3.17 Undiscounted estimate of effects from intervention based on North
Karelia project

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 3.53 | 0.41 | 26.4 | 38.0 | 62 |
| Males 50-59 | 4.72 | 0.68 | 33.0 | 40.2 | 72 |
| Males 60-69 | 5.57 | 0.86 | 30.5 | 33.4 | 70 |
| Males 70-79 | 5.54 | 0.79 | 18.4 | 19.2 | 54 |
| Females 40-49 | 3.42 | 0.34 | 27.4 | 39.8 | 72 |
| Females 50-59 | 3.84 | 0.53 | 31.5 | 38.2 | 68 |
| Females 60-69 | 4.66 | 0.76 | 32.0 | 34.2 | 68 |
| Females 70-79 | 5.12 | 0.77 | 22.3 | 22.1 | 56 |

Table 3.18 Discounted estimate of effects from intervention based on North Karelia project

| Per 1000 | Cases prevented | Deaths prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 3.05 | 0.35 | 8.4 | 14.3 | 32 |
| Males 50-59 | 4.15 | 0.60 | 13.2 | 18.0 | 45 |
| Males 60-69 | 5.01 | 0.78 | 15.3 | 18.3 | 51 |
| Males 70-79 | 5.06 | 0.72 | 11.5 | 12.8 | 45 |
| Females 40-49 | 2.93 | 0.29 | 7.8 | 13.8 | 35 |
| Females 50-59 | 3.32 | 0.45 | 11.3 | 15.8 | 40 |
| Females 60-69 | 4.09 | 0.67 | 14.8 | 17.4 | 47 |
| Females 70-79 | 4.59 | 0.69 | 13.1 | 13.8 | 45 |

Table 3.19 Undiscounted estimates of total population effects from intervention based on North Karelia project

|  | 000s of cases prevented | 000s of deaths prevented | $\begin{aligned} & \text { 000s of LY } \\ & \text { gained } \end{aligned}$ | 000s of QALYs gained | £millions saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 14 | 1.6 | 103 | 148 | 240 |
| Males 50-59 | 15 | 2.2 | 105 | 128 | 230 |
| Males 60-69 | 14 | 2.2 | 78 | 86 | 180 |
| Males 70-79 | 9 | 1.3 | 31 | 32 | 91 |
| Females 40-49 | 14 | 1.3 | 109 | 159 | 287 |
| Females 50-59 | 13 | 1.7 | 104 | 126 | 226 |
| Females 60-69 | 13 | 2.1 | 88 | 94 | 187 |
| Females 70-79 | 10 | 1.6 | 45 | 45 | 114 |
| Totals | 102 | 14.0 | 664 | 818 | 1555 |

Table 3.20 Discounted estimates of total population effects from intervention based on North Karelia project

|  | 000s of cases <br> prevented | 000s of deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 12 | 1.4 | 33 | 56 | 126 |
| Males 50-59 | 13 | 1.9 | 42 | 57 | 143 |
| Males 60-69 | 13 | 2.0 | 39 | 47 | 131 |
| Males 70-79 | 8 | 1.2 | 19 | 21 | 75 |
| Females 40-49 | 12 | 1.1 | 31 | 55 | 140 |
| Females 50-59 | 11 | 1.5 | 37 | 52 | 133 |
| Females 60-69 | 11 | 1.8 | 41 | 48 | 130 |
| Females 70-79 | 9 | 1.4 | 26 | 28 | 91 |
| Totals | 90 | 12.3 | 269 | 364 | 969 |

For the cost-effectiveness analysis of this intervention, Tosteson et al (1997) report that it was costed at US $\$ 10$ (price year 1985) per person reached for the first year, and US $\$ 5$ per year thereafter. Converting into sterling and inflating to 2008 gives us a cost estimate of $£ 30$ per person for the first year and $£ 15$ thereafter. Applying these costs over ten years gives us an estimated (discounted) cost of $£ 144$ per person reached.

This should be multiplied by the total population in the age range 40-79 to give us an estimated project cost of $£ 144 \times 24,000,000=£ 3.5$ billion. Table 3.20 shows a saving in healthcare costs of approximately $£ 1$ billion. This gives us a net cost of $£ 2.5$ billion for a gain of approximately 360,000 QALY at an incremental cost-effectiveness ratio of approximately $£ 7,000 /$ QALY. This is without taking into account the benefit from smoking reduction.

### 3.2.2 The Stanford Five City project

Tosteson et al (1997) report results for the Stanford Five City Project of a 4\% reduction in systolic blood pressure and $2 \%$ decrease in serum cholesterol, achieved at a cost of $\$ 4.95$ per person per year (price year 1993). Converting to sterling and inflating to 2008 gives a cost of $£ 5.05$ per person per year. A discounted total over 10 years is then about $£ 44$, which multiplies up to about $£ 1$ billion total cost for the project. Tables 3.21 to 3.24 show the outcomes from the "Risk Equation Modifying" model. The total healthcare cost saving almost equals the estimated cost of the project and it is not sensible to quote an ICER in this case. Including any appreciable benefit from smoking reduction would make the programme cost saving as well as improving health.

Table 3.21 Undiscounted estimate of effects from intervention based on Stanford Five
City project

| Per 1000 | Cases prevented | Deaths prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 3.68 | 0.43 | 27.5 | 39.5 | 64 |
| Males 50-59 | 4.91 | 0.71 | 34.4 | 41.9 | 75 |
| Males 60-69 | 5.80 | 0.90 | 31.7 | 34.8 | 73 |
| Males 70-79 | 5.78 | 0.82 | 19.1 | 20.0 | 57 |
| Females 40-49 | 3.28 | 0.32 | 26.3 | 38.2 | 69 |
| Females 50-59 | 3.68 | 0.50 | 30.2 | 36.6 | 65 |
| Females 60-69 | 4.46 | 0.73 | 30.7 | 32.8 | 65 |
| Females 70-79 | 4.91 | 0.74 | 21.4 | 21.2 | 54 |

Table 3.22 Discounted estimate of effects from intervention based on Stanford Five City project

| Per 1000 | Cases prevented | Deaths prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 3.17 | 0.37 | 8.8 | 14.9 | 34 |
| Males 50-59 | 4.32 | 0.62 | 13.7 | 18.8 | 47 |
| Males 60-69 | 5.22 | 0.81 | 16.0 | 19.0 | 53 |
| Males 70-79 | 5.27 | 0.75 | 12.0 | 13.3 | 47 |
| Females 40-49 | 2.81 | 0.28 | 7.5 | 13.3 | 34 |
| Females 50-59 | 3.18 | 0.43 | 10.8 | 15.1 | 38 |
| Females 60-69 | 3.92 | 0.64 | 14.2 | 16.7 | 45 |
| Females 70-79 | 4.40 | 0.66 | 12.5 | 13.3 | 43 |

Table 3.23 Undiscounted estimates of total population effects from intervention based on Stanford Five City project

|  | 000s of cases prevented | 000s of deaths prevented | $\begin{aligned} & \text { 000s of LY } \\ & \text { gained } \end{aligned}$ | $\begin{aligned} & \text { 000s of } \\ & \text { QALYs } \\ & \text { gained } \end{aligned}$ | £millions saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 14 | 1.7 | 107 | 154 | 250 |
| Males 50-59 | 16 | 2.3 | 109 | 133 | 240 |
| Males 60-69 | 15 | 2.3 | 81 | 89 | 188 |
| Males 70-79 | 10 | 1.4 | 32 | 33 | 94 |
| Females 40-49 | 13 | 1.3 | 105 | 152 | 275 |
| Females 50-59 | 12 | 1.7 | 100 | 121 | 217 |
| Females 60-69 | 12 | 2.0 | 85 | 90 | 179 |
| Females 70-79 | 10 | 1.5 | 43 | 43 | 109 |
| Totals | 102 | 14.1 | 662 | 817 | 1552 |

Table 3.24 Discounted estimates of total population effects from intervention based on Stanford Five City project

|  | 000s of cases <br> prevented | 000s of deaths <br> prevented | 000s of LY <br> gained | 000 s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 12 | 1.4 | 34 | 58 | 131 |
| Males 50-59 | 14 | 2.0 | 44 | 60 | 148 |
| Males 60-69 | 13 | 2.1 | 41 | 49 | 136 |
| Males 70-79 | 9 | 1.2 | 20 | 22 | 78 |
| Females 40-49 | 11 | 1.1 | 30 | 53 | 134 |
| Females 50-59 | 11 | 1.4 | 36 | 50 | 127 |
| Females 60-69 | 11 | 1.8 | 39 | 46 | 125 |
| Females 70-79 | 9 | 1.3 | 25 | 27 | 87 |
| Totals | 90 | 12.3 | 269 | 364 | 968 |

### 3.2.3 Legislation to ban transfats

The expert paper submitted to the Programme Development Group (Lincoln, 2009) suggests that it is possible to reduce trans fatty acid (TFA) levels by approximately $0.7 \%$ of total fat content, as a conservative estimate, and that a $1 \%$ increase in energy intake from TFAs carries a relative risk of 1.12 of CHD death. For modelling purposes, this relative risk is taken to apply to all CVD events, and so a reduction of $0.7 \%$ gives us a relative risk of $1.12^{-0.7}=0.924$. Putting this figure into the "Relative Risk" model leads to the outcomes shown in Tables 3.25 to 3.28.

Table 3.25 Undiscounted estimate of effects from intervention based on legislation against trans fatty acids

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 6.87 | 0.82 | 51.7 | 73.5 | 120 |
| Males 50-59 | 10.81 | 1.58 | 75.8 | 91.4 | 165 |
| Males 60-69 | 14.62 | 2.24 | 78.2 | 85.4 | 182 |
| Males 70-79 | 15.54 | 2.19 | 49.6 | 51.7 | 148 |
| Females 40-49 | 5.76 | 0.57 | 46.7 | 67.1 | 121 |
| Females 50-59 | 6.98 | 0.98 | 57.7 | 69.3 | 123 |
| Females 60-69 | 9.51 | 1.54 | 64.6 | 68.5 | 137 |
| Females 70-79 | 11.59 | 1.74 | 49.2 | 48.5 | 125 |

Table 3.26 Discounted estimate of effects from intervention based on legislation against trans fatty acids

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 5.85 | 0.69 | 16.4 | 27.4 | 62 |
| Males 50-59 | 9.36 | 1.36 | 30.0 | 40.7 | 101 |
| Males 60-69 | 12.88 | 1.98 | 39.0 | 46.3 | 131 |
| Males 70-79 | 13.86 | 1.96 | 30.7 | 33.9 | 120 |
| Females 40-49 | 4.88 | 0.48 | 13.2 | 23.1 | 59 |
| Females 50-59 | 5.96 | 0.82 | 20.6 | 28.5 | 72 |
| Females 60-69 | 8.23 | 1.34 | 29.8 | 34.6 | 95 |
| Females 70-79 | 10.19 | 1.53 | 28.6 | 30.1 | 99 |

Table 3.27 Undiscounted estimates of total population effects from intervention based on legislation against trans fatty acids

|  | 000s of cases prevented | 000s of deaths prevented | $\begin{aligned} & \text { 000s of LY } \\ & \text { gained } \end{aligned}$ | 000s of QALYs gained | £millions saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 27 | 3.2 | 202 | 286 | 466 |
| Males 50-59 | 34 | 5.0 | 241 | 291 | 524 |
| Males 60-69 | 37 | 5.7 | 200 | 219 | 466 |
| Males 70-79 | 26 | 3.7 | 83 | 86 | 247 |
| Females 40-49 | 23 | 2.3 | 186 | 267 | 480 |
| Females 50-59 | 23 | 3.2 | 191 | 229 | 408 |
| Females 60-69 | 26 | 4.2 | 178 | 189 | 377 |
| Females 70-79 | 23 | 3.5 | 99 | 98 | 252 |
| Totals | 220 | 30.9 | 1380 | 1666 | 3221 |

Table 3.28 Discounted estimates of total population effects from intervention based on legislation against trans fatty acids

|  | 000s of cases <br> prevented | 000s of deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 23 | 2.7 | 64 | 107 | 243 |
| Males 50-59 | 30 | 4.3 | 96 | 129 | 322 |
| Males 60-69 | 33 | 5.1 | 100 | 119 | 335 |
| Males 70-79 | 23 | 3.3 | 51 | 57 | 200 |
| Females 40-49 | 19 | 1.9 | 53 | 92 | 234 |
| Females 50-59 | 20 | 2.7 | 68 | 94 | 239 |
| Females 60-69 | 23 | 3.7 | 82 | 95 | 261 |
| Females 70-79 | 21 | 3.1 | 58 | 61 | 199 |
| Totals | 191 | 26.8 | 571 | 754 | 2033 |

The figures in Table 3.28 suggest that an intervention costing about $£ 2$ billion would still be cost saving if it could achieve the desired effect. Allowing for discounting, the equivalent annual cost would be about $£ 240$ million. An intervention could cost more than this and still be cost-effective. Table 3.29 shows the possibilities here.

Table 3.29 Maximum intervention cost at which an intervention based on legislation against trans fatty acids could be cost-effective

| Threshold ICER <br> $(£ /$ QALY $)$ | Maximum one-off cost | Maximum annual cost <br> over 10 years |
| :---: | :---: | :---: |
| 20,000 | $£ 17$ billion | $£ 2.0$ billion |
| 30,000 | $£ 25$ billion | $£ 2.9$ billion |

The $0.7 \%$ TFA reduction represents a conservative estimate of what is possible. Assuming a uniform decrease from $2 \%$ to $0.5 \%$, in line with effects observed in Denmark, would generate a TFA reduction of $1.5 \%$. Benefits in deprived groups might be substantially larger, given the $6 \%$ daily intake observed in some disadvantaged groups. On the other hand, it is possible that legislation would not achieve the full modelled effect. Section 3.3 of this report gives the maximum acceptable cost for an intervention given a range of possible relative risks.

### 3.2.4 Legislation to reduce salt intake

The expert paper on salt intake (Cappucio, 2009) suggests that a reduction of 3 g per day in salt intake is a reasonably conservative estimate of the potential effects of
legislation to reduce salt intake, and that a reduction of 6 g per day is a reasonable aspiration. According to He and MacGregor (2003), reductions in salt consumption of 3 g per day and 6 g per day would lead to mean reductions in systolic blood pressure of 2.5 mmHg and 5 mmHg respectively. To fit the risk reduction model, these reductions must be expressed as percentages and can be taken as approximately 2 percent and 4 percent respectively. Tables 3.30 to 3.34 give results for a reduction of 3 g per day in salt intake, while Tables 3.35 to 3.39 give equivalents for a reduction of 6 g per day. Summarised results for other values appear in section 3.3.2 where generic interventions to reduce systolic blood pressure are considered.

Table 3.30 Undiscounted estimate of effects from reduction of 3 g per day in salt intake

| Per 1000 | Cases prevented | Deaths prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 1.32 | 0.15 | 9.9 | 14.2 | 23 |
| Males 50-59 | 1.76 | 0.25 | 12.3 | 15.0 | 27 |
| Males 60-69 | 2.08 | 0.32 | 11.3 | 12.5 | 26 |
| Males 70-79 | 2.06 | 0.29 | 6.8 | 7.2 | 20 |
| Females 40-49 | 1.18 | 0.12 | 9.5 | 13.7 | 25 |
| Females 50-59 | 1.32 | 0.18 | 10.8 | 13.2 | 24 |
| Females 60-69 | 1.60 | 0.26 | 11.0 | 11.8 | 23 |
| Females 70-79 | 1.76 | 0.26 | 7.7 | 7.6 | 19 |

Table 3.31 Discounted estimate of effects from reduction of 3 g per day in salt intake

| Per 1000 | Cases <br> prevented | Deaths <br> prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 1.14 | 0.13 | 3.2 | 5.3 | 12 |
| Males 50-59 | 1.55 | 0.22 | 4.9 | 6.7 | 17 |
| Males 60-69 | 1.87 | 0.29 | 5.7 | 6.8 | 19 |
| Males 70-79 | 1.88 | 0.27 | 4.3 | 4.8 | 17 |
| Females 40-49 | 1.01 | 0.10 | 2.7 | 4.8 | 12 |
| Females 50-59 | 1.14 | 0.15 | 3.9 | 5.4 | 14 |
| Females 60-69 | 1.41 | 0.23 | 5.1 | 6.0 | 16 |
| Females 70-79 | 1.57 | 0.24 | 4.5 | 4.8 | 15 |

Table 3.32 Undiscounted estimates of total population effects from reduction of 3 g per day in salt intake

|  | 000s of cases <br> prevented | 000s of <br> deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 5.2 | 0.60 | 39 | 59 | 90 |
| Males 50-59 | 5.6 | 0.81 | 39 | 48 | 86 |
| Males 60-69 | 5.3 | 0.82 | 29 | 32 | 67 |
| Males 70-79 | 3.4 | 0.49 | 11 | 12 | 34 |
| Females 40-49 | 4.7 | 0.46 | 38 | 55 | 99 |
| Females 50-59 | 4.4 | 0.60 | 36 | 44 | 78 |
| Females 60-69 | 4.4 | 0.72 | 30 | 32 | 64 |
| Females 70-79 | 3.5 | 0.53 | 15 | 15 | 39 |
| Totals | 36.6 | 5.04 | 238 | 293 | 557 |

Table 3.33 Discounted estimates of total population effects from reduction of 3 g per day in salt intake

|  | 000s of cases <br> prevented | 000s of deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 4.4 | 0.51 | 12 | 21 | 47 |
| Males 50-59 | 4.9 | 0.71 | 16 | 21 | 53 |
| Males 60-69 | 4.8 | 0.74 | 15 | 17 | 49 |
| Males 70-79 | 3.1 | 0.45 | 7 | 8 | 28 |
| Females 40-49 | 4.0 | 0.39 | 11 | 19 | 48 |
| Females 50-59 | 3.8 | 0.51 | 13 | 18 | 46 |
| Females 60-69 | 3.9 | 0.64 | 14 | 16 | 45 |
| Females 70-79 | 3.2 | 0.48 | 9 | 10 | 31 |
| Totals | 32.2 | 4.43 | 96 | 131 | 347 |

Table 3.34 Maximum intervention cost at which a reduction of 3 g per day in salt intake could be cost-effective

| Threshold ICER <br> $(£ /$ QALY $)$ | Maximum one-off cost | Maximum annual cost <br> over 10 years |
| :---: | :---: | :---: |
| 20,000 | $£ 3.0$ billion | $£ 340$ million |
| 30,000 | $£ 4.3$ billion | $£ 500$ million |

Table 3.35 Undiscounted estimate of effects from reduction of 6 g per day in salt intake

| Per 1000 | Cases prevented | Deaths prevented | LY gain | QALY gain | £000 saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 2.66 | 0.31 | 19.9 | 28.6 | 46 |
| Males 50-59 | 3.55 | 0.51 | 24.8 | 30.2 | 54 |
| Males 60-69 | 4.18 | 0.65 | 22.9 | 25.1 | 53 |
| Males 70-79 | 4.16 | 0.59 | 13.8 | 14.4 | 41 |
| Females 40-49 | 2.37 | 0.23 | 19.0 | 27.6 | 50 |
| Females 50-59 | 2.66 | 0.36 | 21.8 | 26.5 | 47 |
| Females 60-69 | 3.22 | 0.53 | 22.2 | 23.7 | 47 |
| Females 70-79 | 3.54 | 0.53 | 15.4 | 15.3 | 39 |

Table 3.36 Discounted estimate of effects from reduction of 6 g per day in salt intake

|  | Cases | Deaths |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Per 1000 | prevented | prevented | LY gain | QALY gain | £000 saved |
| Males 40-49 | 2.29 | 0.27 | 6.3 | 10.7 | 24 |
| Males 50-59 | 3.12 | 0.45 | 9.9 | 13.6 | 34 |
| Males 60-69 | 3.76 | 0.58 | 11.5 | 13.7 | 38 |
| Males 70-79 | 3.80 | 0.54 | 8.6 | 9.6 | 34 |
| Females 40-49 | 2.03 | 0.20 | 5.4 | 9.6 | 24 |
| Females 50-59 | 2.30 | 0.31 | 7.8 | 10.9 | 28 |
| Females 60-69 | 2.83 | 0.46 | 10.3 | 12.0 | 33 |
| Females 70-79 | 3.17 | 0.48 | 9.0 | 9.6 | 31 |

Table 3.37 Undiscounted estimates of total population effects from reduction of 6 g per day in salt intake

|  | 000s of cases <br> prevented | 000s of <br> deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 10.4 | 1.21 | 77 | 111 | 181 |
| Males 50-59 | 11.3 | 1.63 | 79 | 96 | 173 |
| Males 60-69 | 10.7 | 1.66 | 59 | 64 | 135 |
| Males 70-79 | 7.0 | 0.98 | 23 | 24 | 68 |
| Females 40-49 | 9.4 | 0.93 | 76 | 110 | 199 |
| Females 50-59 | 8.8 | 1.20 | 72 | 88 | 157 |
| Females 60-69 | 8.9 | 1.45 | 61 | 65 | 129 |
| Females 70-79 | 7.1 | 1.07 | 31 | 31 | 79 |
| Totals | 73.6 | 10.15 | 478 | 590 | 1120 |

Table 3.38 Discounted estimates of total population effects from reduction of 6 g per day in salt intake

|  | 000s of cases <br> prevented | 000s of deaths <br> prevented | 000s of LY <br> gained | 000s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Males 40-49 | 8.9 | 1.03 | 25 | 42 | 95 |
| Males 50-59 | 9.9 | 1.42 | 31 | 43 | 107 |
| Males 60-69 | 9.6 | 1.50 | 29 | 35 | 98 |
| Males 70-79 | 6.3 | 0.90 | 14 | 16 | 56 |
| Females 40-49 | 8.1 | 0.79 | 22 | 38 | 97 |
| Females 50-59 | 7.6 | 1.02 | 26 | 36 | 92 |
| Females 60-69 | 7.8 | 1.28 | 28 | 33 | 90 |
| Females 70-79 | 6.4 | 0.96 | 18 | 19 | 63 |
| Totals | 64.8 | 8.91 | 194 | 263 | 699 |

Table 3.39 Maximum intervention cost at which a reduction of 6 g per day in salt intake could be cost-effective

| Threshold ICER <br> $(£ /$ QALY $)$ | Maximum one-off cost | Maximum annual cost <br> over 10 years |
| :---: | :---: | :---: |
| 20,000 | $£ 6.0$ billion | $£ 700$ million |
| 30,000 | $£ 8.6$ billion | $£ 1.0$ billion |

### 3.3 Results for hypothetical interventions

In this section, we consider hypothetical interventions and estimate the maximum acceptable cost of an intervention to achieve a given effect. As in previous sections, this is applied to a population with a starting age of 40-79 and estimates the lifetime effects of events prevented over 10 years.

### 3.3.1 Relative risk

Using the "Relative Risk" model, Table 3.40 shows the estimated outcomes for an intervention achieving a given relative risk of a primary CVD event, assuming that this applies uniformly across the modelled population. The equivalent discounted figures are shown in Table 3.41, and the acceptable costs for cost saving, or staying within a cost-effectiveness threshold of $£ 20,000$ or $£ 30,000$ per QALY appear in Table 3.42. A 0.5 reduction in relative risk, halving the CVD burden, was estimated to generate discounted savings of approximately $£ 14$ billion (Table 3.41).

Table 3.40 Undiscounted outcomes for intervention with given relative risk

|  | $000 s$ <br> cases | 000 s of <br> deaths <br> prevented <br> prevented | 000 s of <br> LY <br> gained | 000 s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0.999 | 3 | 0.4 | 18 | 22 | 42 |
| 0.995 | 14 | 2.0 | 90 | 109 | 210 |
| 0.99 | 29 | 4.0 | 180 | 217 | 420 |
| 0.98 | 57 | 8.1 | 360 | 435 | 841 |
| 0.97 | 86 | 12 | 541 | 653 | 1262 |
| 0.96 | 115 | 16 | 722 | 871 | 1684 |
| 0.95 | 144 | 20 | 903 | 1090 | 2107 |
| 0.94 | 173 | 24 | 1084 | 1309 | 2531 |
| 0.93 | 202 | 28 | 1266 | 1528 | 2955 |
| 0.92 | 231 | 32 | 1448 | 1748 | 3380 |
| 0.91 | 260 | 37 | 1631 | 1968 | 3806 |
| 0.9 | 289 | 41 | 1814 | 2188 | 4232 |
| 0.85 | 436 | 61 | 2732 | 3295 | 6376 |
| 0.8 | 585 | 82 | 3659 | 4412 | 8538 |
| 0.75 | 735 | 103 | 4593 | 5537 | 10718 |
| 0.7 | 886 | 124 | 5536 | 6671 | 12918 |
| 0.65 | 1039 | 146 | 6487 | 7815 | 15136 |
| 0.6 | 1193 | 168 | 7447 | 8967 | 17374 |
| 0.55 | 1349 | 190 | 8414 | 10129 | 19631 |
| 0.5 | 1507 | 212 | 9390 | 11301 | 21909 |

Table 3.41 Discounted outcomes for intervention with given relative risk

|  | 000 s of <br> cases <br> prevented | 000 s of <br> deaths <br> prevented | 000 s of <br> LY <br> gained | 000 of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0.999 | 2 | 0.3 | 7 | 10 | 26 |
| 0.995 | 12 | 1.7 | 37 | 49 | 132 |
| 0.99 | 25 | 3.5 | 74 | 98 | 265 |
| 0.98 | 50 | 7.0 | 149 | 197 | 530 |
| 0.97 | 75 | 10 | 224 | 295 | 796 |
| 0.96 | 100 | 14 | 299 | 394 | 1063 |
| 0.95 | 125 | 18 | 374 | 493 | 1330 |
| 0.94 | 150 | 21 | 449 | 592 | 1597 |
| 0.93 | 175 | 25 | 524 | 692 | 1865 |
| 0.92 | 201 | 28 | 600 | 791 | 2133 |
| 0.91 | 226 | 32 | 675 | 891 | 2402 |
| 0.9 | 251 | 35 | 751 | 990 | 2671 |
| 0.85 | 378 | 53 | 1132 | 1492 | 4024 |
| 0.8 | 507 | 71 | 1516 | 1997 | 5389 |
| 0.75 | 637 | 89 | 1903 | 2507 | 6766 |
| 0.7 | 768 | 108 | 2294 | 3021 | 8155 |
| 0.65 | 900 | 126 | 2689 | 3540 | 9557 |
| 0.6 | 1033 | 145 | 3088 | 4062 | 10971 |
| 0.55 | 1168 | 164 | 3490 | 4589 | 12397 |
| 0.5 | 1304 | 183 | 3895 | 5121 | 13836 |

Table 3.42 Maximum acceptable cost (£million per programme) for intervention with given relative risk

|  | For cost saving |  | For $£ 20,000 /$ QALY |  | For $£ 30,000 / \mathrm{QALY}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Relative risk | one-off | annual | one-off | annual | one-off | annual |
| 0.999 | 26 | 3 | 223 | 26 | 321 | 37 |
| 0.995 | 132 | 15 | 1115 | 130 | 1607 | 187 |
| 0.99 | 265 | 31 | 2231 | 259 | 3214 | 373 |
| 0.98 | 530 | 62 | 4466 | 519 | 6434 | 747 |
| 0.97 | 796 | 93 | 6705 | 779 | 9659 | 1122 |
| 0.96 | 1063 | 123 | 8947 | 1039 | 12890 | 1497 |
| 0.95 | 1330 | 154 | 11193 | 1300 | 16125 | 1873 |
| 0.94 | 1597 | 186 | 13443 | 1562 | 19366 | 2250 |
| 0.93 | 1865 | 217 | 15697 | 1824 | 22613 | 2627 |
| 0.92 | 2133 | 248 | 17954 | 2086 | 25865 | 3005 |
| 0.91 | 2402 | 279 | 20215 | 2349 | 29122 | 3383 |
| 0.9 | 2671 | 310 | 22480 | 2612 | 32385 | 3762 |
| 0.85 | 4024 | 467 | 33862 | 3934 | 48780 | 5667 |
| 0.8 | 5389 | 626 | 45338 | 5267 | 65313 | 7588 |
| 0.75 | 6766 | 786 | 56911 | 6612 | 81984 | 9524 |
| 0.7 | 8155 | 947 | 68582 | 7967 | 98795 | 11477 |
| 0.65 | 9557 | 1110 | 80350 | 9335 | 115747 | 13447 |
| 0.6 | 10971 | 1275 | 92218 | 10713 | 132842 | 15433 |
| 0.55 | 12397 | 1440 | 104187 | 12104 | 150081 | 17436 |
| 0.5 | 13836 | 1607 | 116256 | 13506 | 167466 | 19455 |

### 3.3.2 Percentage reduction in systolic blood pressure

Using the "Risk Equation Modifying" model, Table 3.43 shows the estimated outcomes for an intervention achieving a given percentage reduction in systolic blood pressure, assuming that this applies uniformly across the modelled population.

Table 3.43 Undiscounted outcomes for intervention with given percentage reduction in systolic blood pressure

|  | Equivalent <br> salt | 000s of <br> cases <br> Peduction <br> reduction in SBP | 000s of <br> (g/day) | 00aths <br> prevented | LY <br> prevented <br> gained | 000s of <br> QALYs <br> gained |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | | £millions |
| :---: |
| 0.5 |

The equivalent discounted figures are shown in Table 3.44, and the acceptable costs for cost saving, or staying within a cost-effectiveness threshold of $£ 20,000$ or $£ 30,000$ per QALY appear in Table 3.45. In each of these tables, the equivalent daily reduction in salt intake is also shown, as explained in Section 3.2.4 above.

Table 3.44 Discounted outcomes for intervention with given percentage reduction in systolic blood pressure

| Percentage reduction in SBP | Equivalent salt reduction (g/day) | 000s of cases prevented | 000s of deaths prevented | 000s of LY gained | 000s of <br> QALYs <br> gained | £millions saved |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.5 | 0.75 | 8 | 1.1 | 24 | 33 | 86 |
| 1 | 1.5 | 16 | 2.2 | 48 | 65 | 173 |
| 1.5 | 2.25 | 24 | 3.3 | 72 | 98 | 260 |
| 2 | 3 | 32 | 4.4 | 96 | 131 | 347 |
| 2.5 | 3.75 | 40 | 5.5 | 121 | 164 | 435 |
| 3 | 4.5 | 48 | 6.7 | 145 | 197 | 522 |
| 3.5 | 5.25 | 57 | 7.8 | 169 | 230 | 610 |
| 4 | 6 | 65 | 8.9 | 194 | 263 | 699 |
| 4.5 | 6.75 | 73 | 10.0 | 219 | 296 | 787 |
| 5 | 7.5 | 81 | 11.2 | 243 | 330 | 876 |

Table 3.45 Maximum acceptable cost (£million per programme) for intervention with given percentage reduction in systolic blood pressure

| Percentage reduction in SBP | Equivalent salt reduction (g/day) | For cost saving |  | For |  | For | QALY |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | one-off | annual | one-off | annual | one-off | annual |
| 0.5 | 0.75 | 86 | 10 | 737 | 86 | 1062 | 123 |
| 1 | 1.5 | 173 | 20 | 1476 | 172 | 2128 | 247 |
| 1.5 | 2.25 | 260 | 30 | 2218 | 258 | 3197 | 371 |
| 2 | 3 | 347 | 40 | 2962 | 344 | 4269 | 496 |
| 2.5 | 3.75 | 435 | 50 | 3708 | 431 | 5344 | 621 |
| 3 | 4.5 | 522 | 61 | 4456 | 518 | 6423 | 746 |
| 3.5 | 5.25 | 610 | 71 | 5207 | 605 | 7505 | 872 |
| 4 | 6 | 699 | 81 | 5960 | 692 | 8590 | 998 |
| 4.5 | 6.75 | 787 | 91 | 6715 | 780 | 9679 | 1124 |
| 5 | 7.5 | 876 | 102 | 7473 | 868 | 10771 | 1251 |

### 3.3.3 Percentage reduction in cholesterol

Using the "Risk Equation Modifying" model, Table 3.46 shows the estimated outcomes for an intervention achieving a given percentage reduction in cholesterol, assuming that this applies uniformly across the modelled population. The equivalent discounted figures are shown in Table 3.47, and the acceptable costs for cost saving,
or staying within a cost-effectiveness threshold of $£ 20,000$ or $£ 30,000$ per QALY appear in Table 3.48.

Table 3.46 Undiscounted outcomes for intervention with given percentage reduction in cholesterol

| Percentage <br> reduction in <br> cholesterol | 000 s of <br> cases <br> prevented | 000 s of <br> deaths <br> prevented | 000 s of <br> LY <br> gained | 000 s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0.5 | 7 | 1.0 | 47 | 57 | 109 |
| 1 | 14 | 2.0 | 93 | 115 | 219 |
| 1.5 | 22 | 3.0 | 140 | 173 | 329 |
| 2 | 29 | 4.0 | 187 | 231 | 439 |
| 2.5 | 36 | 5.0 | 234 | 289 | 550 |
| 3 | 43 | 6.0 | 282 | 348 | 661 |
| 3.5 | 51 | 7.0 | 329 | 406 | 772 |
| 4 | 58 | 8.0 | 377 | 465 | 884 |
| 4.5 | 65 | 9.0 | 425 | 524 | 996 |
| 5 | 73 | 10.0 | 473 | 583 | 1109 |

Table 3.47 Discounted outcomes for intervention with given percentage reduction in cholesterol

| Percentage <br> reduction in <br> cholesterol | 000 s of <br> cases <br> prevented | 000 s of <br> deaths <br> prevented | 000 s of <br> LY <br> gained | 000 s of <br> QALYs <br> gained | £millions <br> saved |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0.5 | 6 | 0.9 | 19 | 26 | 68 |
| 1 | 13 | 1.7 | 38 | 51 | 136 |
| 1.5 | 19 | 2.6 | 57 | 77 | 205 |
| 2 | 25 | 3.5 | 76 | 103 | 274 |
| 2.5 | 32 | 4.4 | 95 | 129 | 343 |
| 3 | 38 | 5.3 | 114 | 155 | 412 |
| 3.5 | 45 | 6.1 | 134 | 181 | 481 |
| 4 | 51 | 7.0 | 153 | 208 | 551 |
| 4.5 | 58 | 7.9 | 172 | 234 | 621 |
| 5 | 64 | 8.8 | 192 | 260 | 691 |

Table 3.48 Maximum acceptable cost (£million per programme) for intervention with given percentage reduction in cholesterol

| Percentage <br> reduction in <br> cholesterol | For cost saving <br> one-off <br> annual | For $£ 20,000 /$ QALY <br> one-off <br> one-off |  | For $£ 30,000 /$ QALY <br> annual |  | one-off |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.5 | 68 | 8 | 581 | 67 | 837 | 97 |
| 1 | 136 | 16 | 1163 | 135 | 1676 | 195 |
| 1.5 | 205 | 24 | 1748 | 203 | 2519 | 293 |
| 2 | 274 | 32 | 2334 | 271 | 3365 | 391 |
| 2.5 | 343 | 40 | 2923 | 340 | 4213 | 489 |
| 3 | 412 | 48 | 3514 | 408 | 5065 | 588 |
| 3.5 | 481 | 56 | 4106 | 477 | 5919 | 688 |
| 4 | 551 | 64 | 4701 | 546 | 6776 | 787 |
| 4.5 | 621 | 72 | 5298 | 616 | 7637 | 887 |
| 5 | 691 | 80 | 5897 | 685 | 8501 | 988 |

## 4. Discussion

The results in Chapter 3 strongly suggest that any legislative intervention that is likely to achive an appreciable reduction in risk of CVD can be expected to produce a net cost saving to the public sector as well as improving health. Only if a very large sum of money needs to be spent in implementing the legislation would this cease to be the case. The conclusion that population wide primary CVD prevention is likely to be cost-saving is reassuringly consistent with findings elsewhere (see for example Abelson et al, 2001, Catford, 2009, Trust for America's Health, 2008, Wanless, 2004). Findings are also consistent with a recent FSA report on salt reduction. Their five year campaign cost approximately $£ 15$ million and achieved a reduction of 0.9 g per day, representing approximately 6000 fewer cardiovascular deaths per year.

A spreadsheet model has been developed which will allow a relative risk to be applied to each year's risk of primary CVD within the population. An alternative form allows percentage reductions in cholesterol and systolic blood pressure to be applied separately for males and females. The reductions modelled of up to $5 \%$ for systolic blood pressure and cholesterol are entirely consistent with the reductions carefully documented in regional programmes such as North Karelia, Stanford and HeartBeat Wales. In their current forms, the models have been built on the assumption that these effects apply uniformly across age and risk groups and, in the case of the "Relative Risk" model, across the ten years and equally for males and females. It would not be difficult to amend the model to allow variation in the effect by any of these factors if such amendment were felt appropriate.

A specific example modelled a $0.7 \%$ reduction in daily trans fatty acid (TFA) intake. A uniform decrease from $2 \%$ to $0.5 \%$ as seen in Denmark would generate a TFA reduction twice as large. Further, the benefits in deprived groups might be larger still, given the $6 \%$ daily intake reported in some UK groups.

From a modelling point of view, the nature of the intervention is unimportant in itself. What matters is whether a given outcome can be achieved (by any type of intervention) for a given cost.

### 4.1 Strengths of the analysis

The model is designed to be transparent to the reader, and involves relatively few assumptions, the effect of each of which can be easily tested. The estimates are based on a series of conservative assumptions, so the true benefits are likely to be substantially larger than reported here.

A 0.5 reduction in the relative risk of CVD was estimated to generate discounted savings of approximately $£ 14$ billion. This is consistent with the results of LuengoFernández and colleagues (2006), who, with a very different methodology, estimated the total burden of CVD to cost $£ 29$ billion.

Results have been given for cholesterol and systolic blood pressure reductions of no more than $5 \%$. In fact, larger reductions in entire populations have been documented in recent years. For cholesterol, reductions have been reported of $22 \%$ in Finland since 1972, 14\% in Iceland since 1980, and 10\% in Sweden since 1986 (Laatikainen et al, 2005, Asplund et al, 2009, Björck et al, 2009). For blood pressure, reductions have been reported of $7.7 \%$ in England since 1981, 6.5\% in Finland since 1972 and $3.5 \%$ in Italy since 1980 (Unal et al, 2004, Laatikainen et al, 2005, Palmieri et al, 2009).

When considering multifactor programmes, a conservative feature of the modelling is that the effects of reduction in smoking prevalence have been omitted. These would generate substantial further reductions in mortality and morbidity, with corresponding financial savings (Unal et al, 2004, Laatikainen et al, 2005, Palmieri et al, 2009).

A further conservative feature is the focus on primary prevention alone. The sort of programmes considered in this report would also benefit the 3 to 4 million patients with recognised CVD in the UK (Unal et al, 2005).

The model assumed a uniform distribution of benefit across social groups. In fact, it is well recognised that the more deprived groups experience disproportionately more disease, and thus would enjoy extra gain from population wide risk factor reductions. Taking this into account would again serve to increase our current conservative estimate of reduction in mortality and morbidity.

### 4.2 Limitations of the analysis

In a different context, Box and colleagues (1978) stated "all models are wrong, but some are useful". Our conclusion is clearly subject to a number of important limitations necessitated by the nature of the decision problem itself, and by the limited resources available to produce this report.

The most obvious limitation of the modelling carried out here is the lack of a full sensitivity analysis. Many essential data inputs (such as the distribution of risk factors in the population) were only readily available as point estimates. Although it would be possible to assess the effects of specific changes in these parameters, this would not account for the uncertainties inherent in the various modelling assumptions. Any attempt to reflect the uncertainty in parameter inputs would risk a spurious impression that the full uncertainty in the decision problem had been included. However, the results in Section 3.3 provide a measure of sensitivity analysis on the effectiveness of interventions, but this should be interpreted allowing for the uncertainties elsewhere in the model. Further, such a sensitivity analysis would be likely to emphasise the much larger benefits that would be estimated using a less robustly conservative approach.

Much of the population data was available only in 10-year age bands. This applies particularly in relation to the distribution of types of CVD event among primary cases, and accounts for some irregularities in the tables in Appendix 2. However, when these results were applied over a 10 -year period with a starting age of $45,55,65$, or 75 , the effects of the irregularities would tend to cancel out to a large extent, so that the relative error in the overall figures will be small compared to the relative error in the individual years' figures.

Apart from the increased mortality following a non-fatal primary CVD event, no attempt has been made to include recurrence. This means that the estimates of life years lost and cost savings are likely to be underestimates, and, to that extent, the analysis is somewhat conservative. Additionally, benefits over a lifetime would clearly be greater than those restricted to the 10 year horizon used in the model for events prevented.

No attempt has been made to include the effects of smoking cessation in the analysis of multi-factor interventions using the "Risk Factor Modifying" model. To do so with any attempt at realism would require the proportion of smokers in each of the risk groups used for the modelling. An assumption of uniformity between these groups is unrealistic. Given the large reductions in smoking prevalence seen in most Western countries, it is likely that substantial mortality and morbidity benefits and cost savings could be made as a result of smoking cessation as part of such programmes.

Similarly, the analysis is restricted to effects on primary CVD prevention. An intervention which is recommended on the basis of this analysis and is known to have only beneficial effects on other aspects of health can be recommended more strongly as a result. Benefits have also been restricted to the population without CVD. The four million CVD patients in the UK would also benefit from population wide reductions in cholesterol, blood pressure, and smoking.

The analysis is limited to the effects on people aged between 40 and 79 at the time of the intervention. The main reason for this is that this is the age range for which the distribution of population risks was available. At various points in the analysis, data which properly relates only to ages 70-74 has been applied to ages 70-79. Given the very high event rates in elderly individuals, substantial additional benefits might reasonably be expected.

A further limitation is the 10 year time frame for prevention of cases. In the undiscounted analysis, a case postponed within this 10 year time frame does not contribute to the total reckoning for cases prevented, but the life years and QALYs lost are lower for a case postponed, so there is at least that measure of benefit. A case postponed beyond the 10 year times frame is credited at its full value. The discounted analysis gives a further benefit to postponed cases.

### 4.3 Recommendations for further research

Various groups studying both prevention and treatment of CVD have developed more sophisticated models than the one used here. It would be helpful to decision making bodies if such groups could be encouraged to produce results from their models in a similar format to the tables in Section 3.3 and Appendix 2, and to maintain these
tables as further information becomes available. In particular, they will need updating as necessary to take account of changes in background characteristics including mortality and changes in preferred treatment of primary CVD. This will require adequate resources for the modelling. The sum required will, however, be much smaller than the value gained from the policy decisions that such modelling will support.

Further research into the causal links in the model, in particular the underlying risk equation, is likely to be of some benefit. However, it is unlikely that small changes in such factors will change policy conclusions and such research would mainly serve the purpose of ensuring that the model remains up to date.

### 4.4 Conclusions

Population-wide prevention interventions appear consistently powerful and costsaving. The general consistency with results from very different methodologies in the USA, Australia, and the UK Treasury is reassuring.

This is a relatively simple and transparent model with clearly acknowledged limitations. The cumulative conservative assumptions mean that the scale of current benefits and cost-savings are almost certainly under-estimated.

## 5. Appendices

## Appendix 1. Life expectancy calculations

Suppose an individual has an annual risk of mortality $\lambda$. Then the probability of survival at time $t$ is $e^{-\lambda t}$, so the life expectancy is $\int_{0}^{\infty} e^{-\lambda t} d t=\left[\frac{-1}{\lambda} e^{-\lambda t}\right]_{0}^{\infty}=\frac{1}{\lambda}$. Thus a known life expectancy can be approximated by an annual risk equal to the reciprocal of the life expectancy.

Now suppose the individual has additional mortality $\mu$ in the first year and $v$ in subsequent years. Then the life expectancy can be approximated as

$$
\int_{0}^{1} e^{-\alpha+\mu \bar{I}} d t+\int_{1}^{\infty} e^{-\boldsymbol{\alpha}+\mu} e^{-\boldsymbol{\alpha}+v^{-1}-1} d t=\frac{1-e^{-\alpha+\mu}}{\lambda+\mu}+\frac{e^{-\alpha+\mu}}{\lambda+v}=\frac{\lambda+v+e^{-(+\mu)}(l-v)}{(\lambda+\mu)} .
$$

Now $e^{-(+\mu)} \approx 1-(+\mu)$, so the life expectancy can be approximated by

$$
\begin{aligned}
& \frac{\lambda+v+(-\lambda-\mu)(u-v)}{(1+\mu)}=\frac{\lambda+v+\mu-v-\lambda \mu+\lambda v-\mu^{2}+\mu v}{(1+\mu)} \\
& =\frac{\lambda+\mu+\lambda v+\mu v-\lambda \mu-\mu^{2}}{(a+\mu)}=\frac{(a+\mu)(+v-\mu)}{(1+\mu)(1+v)}=\frac{(+v-\mu)}{(a+v)} .
\end{aligned}
$$

This is the formula used to estimate the life expectancy following a particular CVD event.

For discounted life expectancy, note that applying an annual discount rate of $\rho$ effectively multiplies the value of survival at time $t$ by $e^{-\rho t}$. Thus if the undiscounted life expectancy is $\frac{1}{\lambda}$, then the discounted life expectancy may be approximated by $\int_{0}^{\infty} e^{-\lambda t} e^{-\rho t} d t=\frac{1}{\lambda+\rho}$.

The approximations used in developing these formulae are well within the range of reasonable modelling approximations given the assumptions made in other parts of the model.

## Appendix 2. Estimated outcomes following a first CVD event

Tables A2.1 and A2.2 give estimated results for males and females respectively for the outcomes expected from a primary CVD event at any given age. The explanation of the tables is in Section 3.1.1 of this report.

Table A2.1 Estimated outcomes for males experiencing a primary CVD event

|  |  | Undiscounted |  |  | Discounted |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Deaths | Cost | LY lost | QALY lost | Cost | LY lost | QALY lost |
| 40 | 0.101 | 19706 | 8.6372 | 13.7326 | 10629 | 2.6403 | 9.6878 |
| 41 | 0.101 | 19351 | 8.3430 | 13.2795 | 10552 | 2.6048 | 9.5464 |
| 42 | 0.101 | 18994 | 8.0523 | 12.8326 | 10473 | 2.5685 | 9.4031 |
| 43 | 0.101 | 18639 | 7.7683 | 12.3958 | 10392 | 2.5317 | 9.2589 |
| 44 | 0.101 | 18281 | 7.4879 | 11.9651 | 10309 | 2.4940 | 9.1127 |
| 45 | 0.101 | 17921 | 7.2112 | 11.5404 | 10223 | 2.4554 | 8.9644 |
| 46 | 0.101 | 17563 | 6.9410 | 11.1255 | 10135 | 2.4163 | 8.8150 |
| 47 | 0.101 | 17202 | 6.6744 | 10.7166 | 10045 | 2.3763 | 8.6635 |
| 48 | 0.101 | 16844 | 6.4144 | 10.3174 | 9953 | 2.3357 | 8.5108 |
| 49 | 0.101 | 16487 | 6.1608 | 9.9277 | 9859 | 2.2945 | 8.3570 |
| 50 | 0.134 | 18404 | 9.0444 | 11.6798 | 11512 | 3.3684 | 7.4314 |
| 51 | 0.134 | 18037 | 8.6665 | 11.2039 | 11398 | 3.3018 | 7.2940 |
| 52 | 0.134 | 17668 | 8.2980 | 10.7402 | 11280 | 3.2344 | 7.1554 |
| 53 | 0.134 | 17295 | 7.9347 | 10.2840 | 11159 | 3.1653 | 7.0144 |
| 54 | 0.134 | 16922 | 7.5809 | 9.8398 | 11034 | 3.0954 | 6.8720 |
| 55 | 0.134 | 16544 | 7.2325 | 9.4031 | 10905 | 3.0238 | 6.7270 |
| 56 | 0.134 | 16170 | 6.8973 | 8.9825 | 10775 | 2.9522 | 6.5818 |
| 57 | 0.134 | 15791 | 6.5674 | 8.5691 | 10640 | 2.8788 | 6.4339 |
| 58 | 0.134 | 15412 | 6.2467 | 8.1671 | 10501 | 2.8045 | 6.2845 |
| 59 | 0.134 | 15029 | 5.9314 | 7.7723 | 10358 | 2.7284 | 6.1321 |
| 60 | 0.160 | 15226 | 8.3164 | 9.1996 | 11297 | 3.8560 | 5.1415 |
| 61 | 0.160 | 14900 | 7.8775 | 8.7258 | 11151 | 3.7397 | 5.0158 |
| 62 | 0.160 | 14579 | 7.4611 | 8.2759 | 11005 | 3.6250 | 4.8909 |
| 63 | 0.160 | 14250 | 7.0515 | 7.8344 | 10853 | 3.5076 | 4.7632 |
| 64 | 0.160 | 13923 | 6.6588 | 7.4111 | 10698 | 3.3907 | 4.6352 |
| 65 | 0.160 | 13593 | 6.2778 | 7.0007 | 10539 | 3.2727 | 4.5055 |
| 66 | 0.160 | 13265 | 5.9134 | 6.6079 | 10377 | 3.1554 | 4.3755 |
| 67 | 0.160 | 12930 | 5.5556 | 6.2228 | 10209 | 3.0356 | 4.2425 |
| 68 | 0.160 | 12596 | 5.2140 | 5.8548 | 10038 | 2.9166 | 4.1093 |
| 69 | 0.160 | 12260 | 4.8837 | 5.4988 | 9862 | 2.7969 | 3.9743 |
| 70 | 0.143 | 12152 | 5.5870 | 5.8086 | 10310 | 3.2301 | 3.4617 |
| 71 | 0.143 | 11873 | 5.2089 | 5.4314 | 10142 | 3.0812 | 3.3425 |
| 72 | 0.143 | 11596 | 4.8540 | 5.0767 | 9973 | 2.9361 | 3.2246 |
| 73 | 0.143 | 11315 | 4.5114 | 4.7344 | 9798 | 2.7907 | 3.1050 |
| 74 | 0.143 | 11034 | 4.1861 | 4.4089 | 9620 | 2.6474 | 2.9853 |
| 75 | 0.143 | 10757 | 3.8828 | 4.1045 | 9442 | 2.5089 | 2.8675 |
| 76 | 0.143 | 10477 | 3.5909 | 3.8113 | 9258 | 2.3708 | 2.7482 |
| 77 | 0.143 | 10197 | 3.3153 | 3.5338 | 9072 | 2.2356 | 2.6294 |
| 78 | 0.143 | 9924 | 3.0601 | 3.2758 | 8888 | 2.1061 | 2.5131 |
| 79 | 0.143 | 9653 | 2.8200 | 3.0323 | 8702 | 1.9801 | 2.3978 |
| 80 | 0.137 | 8372 | 3.2050 | 3.1524 | 7808 | 2.3017 | 2.0231 |


|  | Undiscounted |  |  |  | Discounted |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Deaths | Cost | LY lost | QALY lost | Cost | LY lost | QALY lost |
| 81 | 0.137 | 8193 | 2.9583 | 2.9175 | 7673 | 2.1622 | 1.9269 |
| 82 | 0.137 | 8015 | 2.7267 | 2.6966 | 7537 | 2.0274 | 1.8320 |
| 83 | 0.137 | 7843 | 2.5149 | 2.4937 | 7404 | 1.9005 | 1.7408 |
| 84 | 0.137 | 7673 | 2.3172 | 2.3040 | 7271 | 1.7789 | 1.6517 |
| 85 | 0.137 | 7510 | 2.1380 | 2.1311 | 7142 | 1.6657 | 1.5668 |
| 86 | 0.137 | 7346 | 1.9672 | 1.9662 | 7012 | 1.5551 | 1.4824 |
| 87 | 0.137 | 7176 | 1.8001 | 1.8050 | 6875 | 1.4442 | 1.3965 |
| 88 | 0.137 | 7006 | 1.6412 | 1.6514 | 6736 | 1.3360 | 1.3110 |
| 89 | 0.137 | 6830 | 1.4862 | 1.5014 | 6590 | 1.2278 | 1.2239 |
| 90 | 0.137 | 6700 | 1.3778 | 1.3939 | 6482 | 1.1504 | 1.1586 |

Table A2.2 Estimated outcomes for females experiencing a primary CVD event

|  | Undiscounted |  |  |  |  | Discounted |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Deaths | Cost | LY lost | QALY lost | Cost | LY lost | QALY lost |  |
| 40 | 0.091 | 24973 | 8.9440 | 14.2751 | 12359 | 2.5067 | 5.0150 |  |
| 41 | 0.091 | 24534 | 8.6544 | 13.8265 | 12271 | 2.4764 | 4.9400 |  |
| 42 | 0.091 | 24092 | 8.3682 | 13.3838 | 12181 | 2.4454 | 4.8641 |  |
| 43 | 0.091 | 23652 | 8.0885 | 12.9510 | 12090 | 2.4140 | 4.7880 |  |
| 44 | 0.091 | 23209 | 7.8121 | 12.5241 | 11995 | 2.3818 | 4.7110 |  |
| 45 | 0.091 | 22763 | 7.5392 | 12.1031 | 11898 | 2.3489 | 4.6332 |  |
| 46 | 0.091 | 22320 | 7.2726 | 11.6917 | 11799 | 2.3156 | 4.5550 |  |
| 47 | 0.091 | 21868 | 7.0067 | 11.2824 | 11696 | 2.2811 | 4.4752 |  |
| 48 | 0.091 | 21423 | 6.7498 | 10.8862 | 11592 | 2.2466 | 4.3957 |  |
| 49 | 0.091 | 20976 | 6.4964 | 10.4958 | 11485 | 2.2112 | 4.3153 |  |
| 50 | 0.106 | 21101 | 9.7704 | 12.9985 | 12581 | 3.2153 | 5.2003 |  |
| 51 | 0.106 | 20708 | 9.3773 | 12.4964 | 12467 | 3.1570 | 5.0959 |  |
| 52 | 0.106 | 20315 | 8.9937 | 12.0069 | 12350 | 3.0980 | 4.9911 |  |
| 53 | 0.106 | 19917 | 8.6157 | 11.5252 | 12230 | 3.0376 | 4.8849 |  |
| 54 | 0.106 | 19514 | 8.2431 | 11.0512 | 12105 | 2.9757 | 4.7773 |  |
| 55 | 0.106 | 19110 | 7.8802 | 10.5895 | 11977 | 2.9131 | 4.6692 |  |
| 56 | 0.106 | 18700 | 7.5229 | 10.1354 | 11845 | 2.8490 | 4.5597 |  |
| 57 | 0.106 | 18295 | 7.1791 | 9.6978 | 11712 | 2.7848 | 4.4506 |  |
| 58 | 0.106 | 17879 | 6.8370 | 9.2633 | 11572 | 2.7184 | 4.3390 |  |
| 59 | 0.106 | 17463 | 6.5045 | 8.8407 | 11428 | 2.6511 | 4.2268 |  |
| 60 | 0.171 | 17915 | 10.2543 | 11.2387 | 12853 | 4.4055 | 5.4347 |  |
| 61 | 0.171 | 17564 | 9.7532 | 10.7006 | 12705 | 4.2890 | 5.2815 |  |
| 62 | 0.171 | 17212 | 9.2705 | 10.1825 | 12554 | 4.1722 | 5.1291 |  |
| 63 | 0.171 | 16853 | 8.7946 | 9.6733 | 12397 | 4.0527 | 4.9744 |  |
| 64 | 0.171 | 16493 | 8.3369 | 9.1836 | 12237 | 3.9330 | 4.8205 |  |
| 65 | 0.171 | 16125 | 7.8862 | 8.7026 | 12070 | 3.8105 | 4.6643 |  |
| 66 | 0.171 | 15754 | 7.4481 | 8.2355 | 11898 | 3.6866 | 4.5075 |  |
| 67 | 0.171 | 15377 | 7.0226 | 7.7821 | 11721 | 3.5612 | 4.3500 |  |
| 68 | 0.171 | 14997 | 6.6095 | 7.3423 | 11538 | 3.4345 | 4.1919 |  |
| 69 | 0.171 | 14613 | 6.2089 | 6.9159 | 11349 | 3.3065 | 4.0331 |  |
| 70 | 0.152 | 13882 | 7.4244 | 7.4120 | 11536 | 4.0009 | 4.3511 |  |
| 71 | 0.152 | 13586 | 6.9456 | 6.9501 | 11364 | 3.8320 | 4.1671 |  |
| 72 | 0.152 | 13288 | 6.4870 | 6.5080 | 11188 | 3.6638 | 3.9848 |  |
| 73 | 0.152 | 12982 | 6.0425 | 6.0800 | 11005 | 3.4943 | 3.8022 |  |
| 74 | 0.152 | 12674 | 5.6180 | 5.6713 | 10818 | 3.3261 | 3.6216 |  |
| 75 | 0.152 | 12363 | 5.2133 | 5.2815 | 10625 | 3.1593 | 3.4433 |  |
| 76 | 0.152 | 12050 | 4.8280 | 4.9100 | 10429 | 2.9944 | 3.2674 |  |
|  |  |  |  |  |  |  |  |  |
| 102 |  |  |  |  |  |  |  |  |


|  | Undiscounted |  |  |  | Discounted |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Deaths | Cost | LY lost | QALY lost | Cost | LY lost | QALY lost |
| 77 | 0.152 | 11735 | 4.4618 | 4.5566 | 10228 | 2.8316 | 3.0942 |
| 78 | 0.152 | 11419 | 4.1144 | 4.2208 | 10022 | 2.6712 | 2.9239 |
| 79 | 0.152 | 11102 | 3.7856 | 3.9022 | 9813 | 2.5135 | 2.7569 |
| 80 | 0.147 | 9675 | 4.4439 | 4.0962 | 8925 | 3.0384 | 2.9440 |
| 81 | 0.147 | 9466 | 4.0876 | 3.7784 | 8770 | 2.8511 | 2.7652 |
| 82 | 0.147 | 9254 | 3.7506 | 3.4777 | 8612 | 2.6679 | 2.5909 |
| 83 | 0.147 | 9040 | 3.4325 | 3.1938 | 8449 | 2.4893 | 2.4212 |
| 84 | 0.147 | 8833 | 3.1446 | 2.9359 | 8290 | 2.3224 | 2.2624 |
| 85 | 0.147 | 8621 | 2.8687 | 2.6887 | 8125 | 2.1576 | 2.1060 |
| 86 | 0.147 | 8413 | 2.6159 | 2.4615 | 7962 | 2.0020 | 1.9581 |
| 87 | 0.147 | 8201 | 2.3746 | 2.2444 | 7793 | 1.8491 | 1.8130 |
| 88 | 0.147 | 7990 | 2.1500 | 2.0417 | 7623 | 1.7026 | 1.6739 |
| 89 | 0.147 | 7780 | 1.9415 | 1.8529 | 7452 | 1.5628 | 1.5411 |
| 90 | 0.147 | 7595 | 1.7688 | 1.6946 | 7300 | 1.4440 | 1.4269 |

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